

Expected Information and its Applications

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Motivation

$\ln(1/p) = -\ln p$ gives the information in
 the **probability** p , $0 \leq p \leq 1$

about the origin

$-\ln(x)$ gives the information in
 the **observation** x , $x \in (0, \infty)$

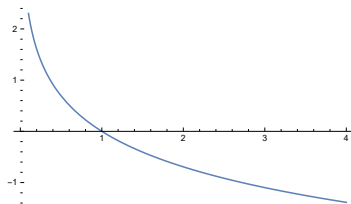


Figure: The function $-\ln x$ on the non-negative reals

Motivation

Given a pdf $f(x)$ on $(0, \infty)$, the function $(-\ln x) \cdot f(x)$ indicates how much information about the origin is contained in the observation x

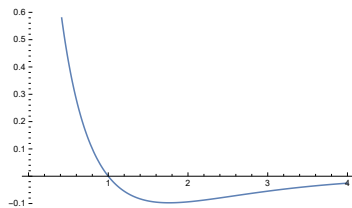
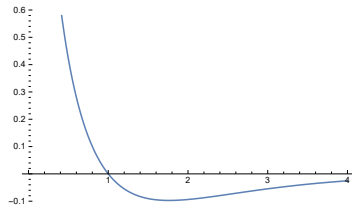


Figure: The function $(-\ln x)e^{-x}$ for the Standard Exponential

Motivation



The overall information in a Standard Exponential is positive, since

$$\int_0^{\infty} (-\ln x) e^{-x} dx = \gamma$$

where $\gamma \approx 0.5772$ is Euler's constant

Definition: Expected Information

Given a pdf $f(x)$, the corresponding expected information is

$$E_I(X) = \int_0^{\infty} (-\ln x) \cdot f(x) dx$$

$$E_I(X) = \int_{-\infty}^{\infty} (-\ln |x|) \cdot f(x) dx$$

$$E_I(X) = \int_{x \in \mathbb{R}^n} (-\ln ||x||) \cdot f(x) dx$$

Only the distance from the origin matters, not the direction

Center vs. Periphery: positive vs. negative information

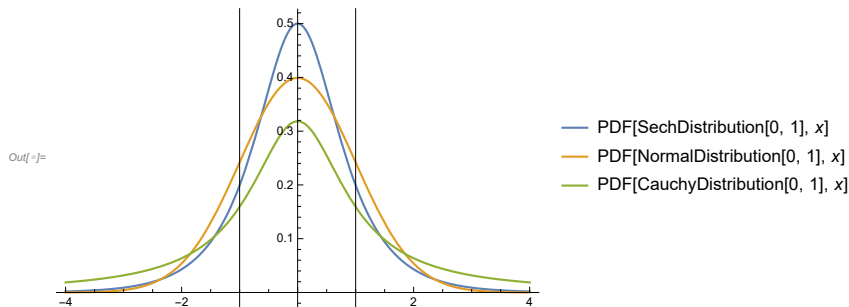
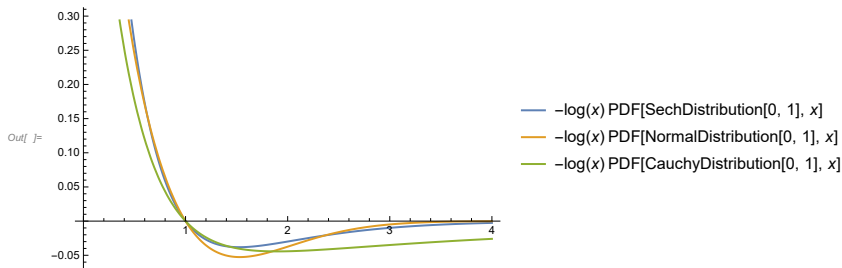


Figure: Center ($|x| < 1$) and Periphery ($|x| > 1$) both contribute to Expected Information

Comparison



$$E_I(\text{Sech}(0, 1)) = 2 \ln \varpi - \ln \pi \approx 0.783$$

$$E_I(\text{Normal}(0, 1)) = (\gamma + \ln 2)/2 \approx 0.635$$

$$E_I(\text{Cauchy}(0, 1)) = 0$$

with the lemniscate constant $\varpi \approx 2.622$

Some Examples

Continuous Distribution	Expected Inform.	Remarks
Laplace(1)	γ	positive
$N \sim \text{Normal}(0, 1)$	$(\gamma + \ln 2)/2$	$> \gamma$
Lévy(0, 1)	$-\gamma - \ln 2$	$1/N^2$
'Lognormal'(0, 1)	0	exp(Normal)
$U \sim \text{Uniform}(0, 1)$	1	elementary
Pareto(1, ∞)	-1	$1/U$
$C \sim \text{Cauchy}(0, 1)$	0	symmetries!
Weibull($a, 1$)	$-\gamma/a$	
Lindley(b)	$\gamma - \frac{1}{b+1} + \ln b$	
Logistic(0, 1)	$\gamma + \ln 2 - \ln \pi$	
Sech(0, 1)	$2 \ln \varpi - \ln \pi$	pdf = $\frac{1}{e^{x\pi/2} + e^{-x\pi/2}}$

$E_I(\cdot)$ highlights how classical distributions are related.

Basic Properties

Whenever expected information exists,

$$E_I(X \cdot Y) = E_I(X) + E_I(Y)$$

In particular

$$E_I(r \cdot X) = E_I(X) + E_I(r) = E_I(X) - \ln |r|$$

Moreover,

$$E_I(X^s) = s \cdot E_I(X),$$

and $E_I(1/X) = -E_I(X)$ in particular.

Exponential mapping: If $X = \exp(-Y)$ then

$$E_I(X) = EY$$

Constructing the Normal: exponential component

Distribution	E_I	Domain	Remark
$U(0, 1)$	1	unit interval	- symmetric -
↓ inverse $P \sim \text{Pareto}$	-1	$(1, \infty)$	polarization
↓ square P^2	-2	$(1, \infty)$	
↓ $\ln(\cdot)$ Exponential(1/2)	$\gamma - \ln 2$	\mathbb{R}^+	nontrivial info.

By the way: the median of an Exponential(1/2) is $2 \ln 2$.

Constructing the Normal: circular component

Distribution	E_I	Domain	Remark
$U(0, 1)$	1	unit interval	- symmetric -
↓ $\cdot \pi/2$ $U(0, \pi/2)$	$1 - \ln \pi + \ln 2$	segment	
↓ $\sin(\cdot)$ pdf = $\frac{2}{\pi\sqrt{1-x^2}}$	$\ln 2$	unit interval	
↓ square pdf = $\frac{1}{\pi\sqrt{x(1-x)}}$			polarization
Arcsin(0, 1)	$2 \ln 2$	unit interval	- symmetric -

Constructing the Normal: putting all together

Distribution	E_I	Domain	Remark
$X \sim \text{Exponential}(1/2)$	$\gamma - \ln 2$	\mathbb{R}^+	
$Y \sim \text{Arcsin}(0, 1)$	$2 \ln 2$	$(0, 1)$	
$X \cdot Y$	$\gamma + \ln 2$	\mathbb{R}^+	
\sqrt{XY}	$(\gamma + \ln 2)/2$	\mathbb{R}^+	depolarization
$S \cdot \sqrt{XY}$	$(\gamma + \ln 2)/2$	\mathbb{R}	symmetrization

Actually,

$$Z = \pm \sqrt{XY} = \pm \sqrt{\ln(U_1^{-2}) \cdot \sin^2\left(\frac{\pi}{2} U_2\right)} \sim N(0, 1)$$

where U_i are independent, Standard Uniform

Entropy and Information Functions

Expected Information:

$$E_I(X) = \int_0^{\infty} (-\ln x) \cdot f(x) dx$$

Entropy is a close non-parametric cousin:

$$H(X) = \int_0^{\infty} (-\ln f(x)) \cdot f(x) dx$$

In general - Integral with respect to an information function $g(x)$:

$$G(X) = \int_0^{\infty} (-\ln g(x)) \cdot f(x) dx$$

Integral Transforms

Laplace:

$$L_f(s) = \int_0^{\infty} e^{-sx} \cdot f(x) dx$$

Mellin:

$$M_f(s) = \int_0^{\infty} x^{s-1} \cdot f(x) dx$$

Logarithmic:

$$T_f(s) = \int_0^{\infty} (-\ln x)^s \cdot f(x) dx$$

Logarithmic Moments and Information Variance

Logarithmic Moments:

$$M_l^s(X) = \int_0^{\infty} (-\ln x)^s f(x) dx$$

Information Variance:

$$\begin{aligned}\sigma_l^2(X) &= M_l^2(X) - (E_l(X))^2 \\ &= \int ((-\ln |x|) - E_l(X))^2 f(x) dx\end{aligned}$$

Exponential mapping: If $X = \exp(-Y)$ then

$$\sigma_l^2(X) = \sigma^2(Y)$$

Properties of Information Variance

Given independence,

$$\sigma_I^2(X \cdot Y) = \sigma_I^2(X) + \sigma^2(Y)$$

$$\sigma^2(X + Y) = \sigma^2(X) + \sigma^2(Y)$$

Further Analogies:

$$\sigma_I^2(a \cdot X) = \sigma_I^2(X) \quad \sigma_I^2(X^b) = b^2 \cdot \sigma_I^2(X)$$

$$\sigma^2(a + X) = \sigma^2(X) \quad \sigma^2(b \cdot X) = b^2 \cdot \sigma^2(X)$$

In particular: $\sigma_I^2(1/X) = \sigma_I^2(X)$

Some Examples

Distribution	E_I	σ_I^2
Normal(0, 1)	$(\gamma + \ln 2)/2$	$\pi^2/8$
Laplace(1)	γ	$\pi^2/6 = 0.1\bar{6}\pi^2$
Cauchy(0, 1)	0	$\pi^2/4 = 0.25\pi^2$
Lévy(0, 1)	$-\gamma - \ln 2$	$\pi^2/2$
Holtzmark	$\gamma/3$	$17\pi^2/108 \approx 0.157\pi^2$
Stable($\alpha, 0$)	$\gamma(1 - 1/\alpha)$	$(1 + 2/\alpha^2)\pi^2/12$
Student t_n	$\frac{\gamma + \ln 2}{2} + \frac{\psi(n/2) - \ln(n/2)}{2}$	$\pi^2/8 + (\psi^{(1)}(n/2))/4$
Gamma(a, b, c)	$-\psi(a)/c - \ln b$	$\psi^{(1)}(a)/c^2$

where the pdf of a Gamma(a, b, c) is

$$f(x) = \frac{c}{b\Gamma(a)} (x/b)^{ac-1} \exp(-(x/b)^c),$$

with the gamma function $\Gamma(x)$, the digamma function

$$\psi(x) = (\ln \Gamma(x))', \text{ and the trigamma function } \psi^{(1)}(x) = \psi'(x)$$

Convergence of Student's t towards the Normal

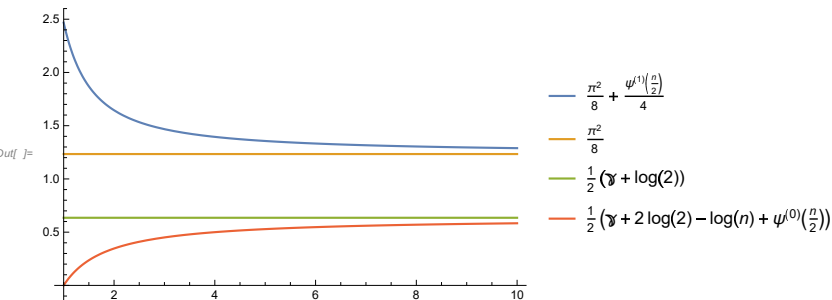


Figure: Expected information and information variance of t_n .

Application: Statistics

Since expected information and information variances add up,

⇒ **Analysis of Information Variance (ANOIVA)**

based on $E_I(XY) = E_I(X) + E_I(Y)$ and $\sigma_I^2(XY) = \sigma_I^2(X) + \sigma_I^2(Y)$

Example: Products of (independent) Normals and Cauchys

	$N(0, 1)$	$N_1 N_2$	N^{-2}	$C(0, 1)$	$C_1 C_2$	C^2
$E(\cdot)$	0	0	–	–	–	–
$\sigma^2(\cdot)$	1	1	–	–	–	–
$E_I(\cdot)$	$(\gamma + \ln 2)/2$	$\gamma + \ln 2$	$-\gamma - \ln 2$	0	0	0
$\sigma_I^2(\cdot)$	$\pi^2/8$	$\pi^2/4$	$\pi^2/2$	$\pi^2/4$	$\pi^2/2$	π^2

The Normal and some of its Cousins

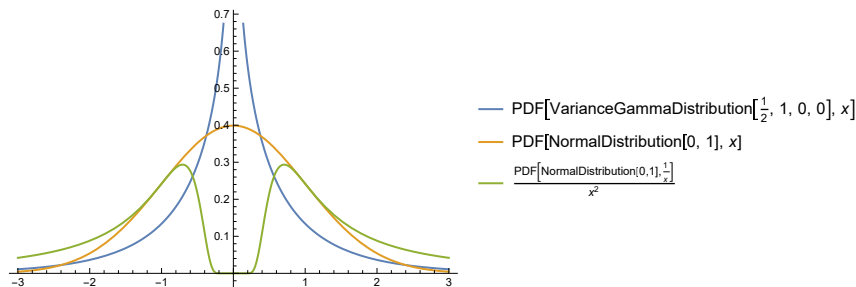


Figure: Product of two Normals, the Standard Normal and its Inverse

More general idea: consider sums, products and powers of random variables, i.e., the **Algebra of Distributions**

Major difference: There are many distributions \mathcal{D} with $E_I(\mathcal{D}) = 0$

A Map of Distributions: the Stochastic Askey Scheme

Notation:

- OPS = Orthogonal Polynomial System
- OPS = Sheffer OPS = Appell and Binomial OPS
- (generalized) hypergeometric function

$${}_rF_s \left(\begin{matrix} a_1, \dots, a_r \\ b_1, \dots, b_s \end{matrix} ; z \right) = \sum_{k=0}^{\infty} \frac{(a_1, \dots, a_r)_k}{(b_1, \dots, b_s)_k} \frac{z^k}{k!}$$

where $(a_1, \dots, a_r)_k = (a_1)_k \cdots (a_r)_k$,

and $(a)_k = \prod_{i=1}^k (a + i - 1)$

is the Pochhammer symbol (rising factorial)

OPS and their corresponding Distributions

Tier	Hypergeom. Function	Orthogonal Poly. System	Distribution (weights)	Convergence
(III)	${}_2F_1(\cdot)$	<i>M.-P.</i>	Sech	\rightarrow (I, IIa)
	${}_2F_1(\cdot)$	Jacobi	Beta	\rightarrow (I, IIa, IIb)
	${}_2F_1(\cdot)$	Pseudo Jacobi	Student t	
	${}_2F_1(\cdot)$	<i>Meixner</i>	Neg. Binom.	\rightarrow (IIa, IIc)
	${}_2F_1(\cdot)$	<i>Krawtchouk</i>	Binomial	\rightarrow (I, IIc)
(IIa)	${}_1F_1(\cdot)$	<i>Laguerre</i>	Gamma	\rightarrow (I)
(IIb)	${}_1F_1(\cdot), {}_2F_0(\cdot)$	Bessel	Inv. Gamma	
(IIc)	${}_2F_0(\cdot)$	<i>Charlier</i>	Poisson	\rightarrow (I)
(I)	${}_2F_0(\cdot)$	<i>Hermite</i>	Normal	

where M.-P. stands for Meixner-Pollaczek

Sheffer OPS \Leftrightarrow Stochastic Askey Scheme

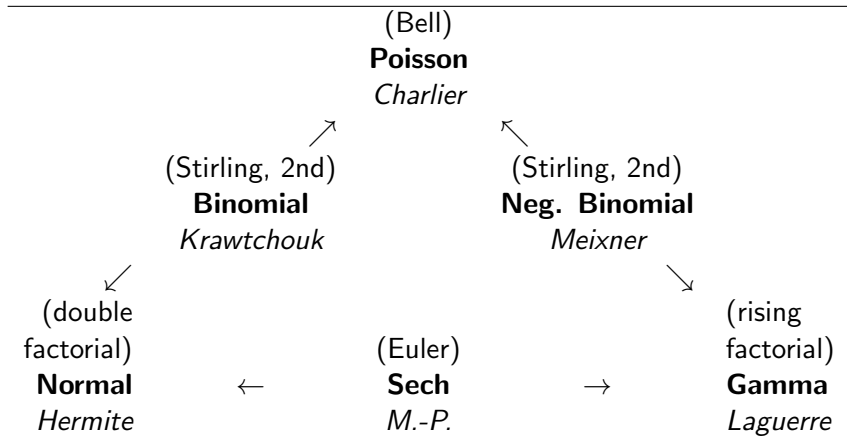
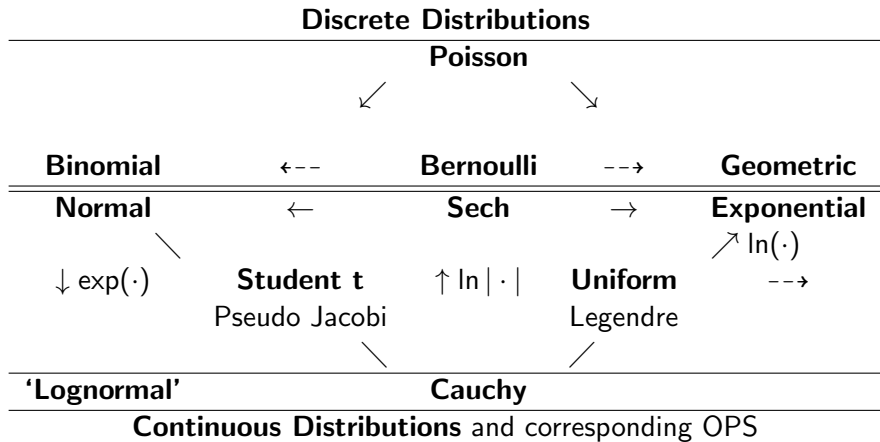


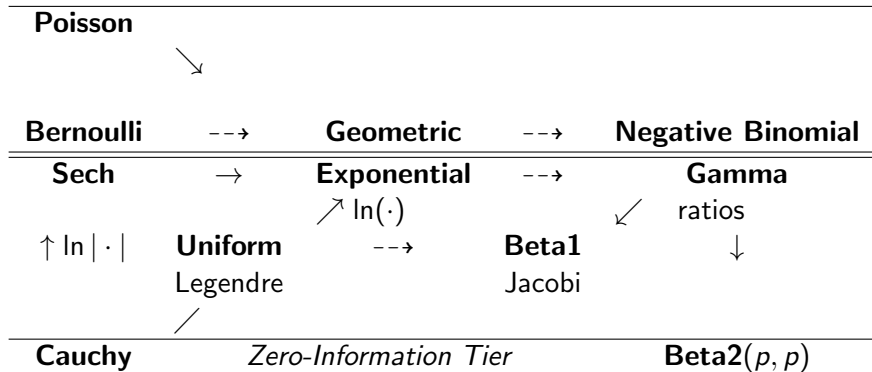
Table: Sheffer OPS, corresp. **Distributions** and (Moments / Numbers)

Extended Stochastic Askey Scheme: Cauchy



Extended Stochastic Askey Scheme: Beta

Discrete Distributions



Continuous Distributions and corresponding OPS

Extended Stochastic Askey Scheme: Below the Line

E_l	Distribution	Ratio Distributions
positive	Gamma <i>Laguerre</i>	Beta1 (p, q) Jacobi
zero		Beta2 (p, p) Jacobi
negative	Inverse Gamma Bessel	Beta1 (q, p) Jacobi

E_l	Product and Ratio Distributions		
positive	Uniform (0, 1) Legendre	Normal N <i>Hermite</i>	VarianceGammaDistrib. $N_1 \cdot N_2$ pdf = $K_0(x)/\pi$
<u>zero</u> ratios:	'NoName' U_1/U_2	Cauchy $C = N_1/N_2$	'Product-Ratio Cauchy' $\frac{N_1 N_2}{N_3 N_4}$ pdf = $\frac{\ln(x^2)}{\pi^2(x^2-1)}$
negative	Pareto OPS = ?	InvNormal $1/N$	Inverse VarianceGamma pdf = $K_0(1/x)/(\pi x^2)$

Logarithmic Transform of Ratio Distributions

Distribution	Ratio	Difference
Laplace	$\ln(U_1/U_2)$	Exponential ₁ – Exponential ₂
Logistic	$\ln(\text{Exp}_1/\text{Exp}_2)$	Gumbel ₁ – Gumbel ₂
Sech	$\ln C $	$\ln N_1 - \ln N_2 $
Csch	$\ln C_1 C_2 $	$\ln N_1 N_3 - \ln N_2 N_4 $
'Morris Family'	$\ln C_1 \cdots C_n $...

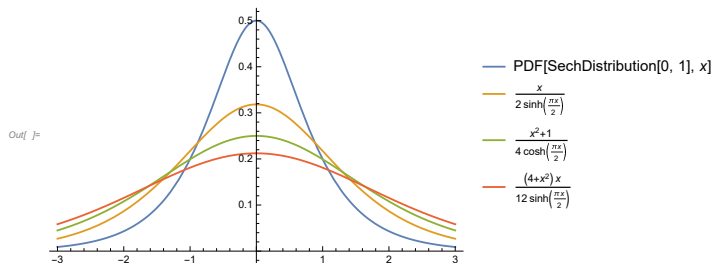


Figure: Generalized Hyperbolic Secant Distributions ('Morris Family')

Interesting Facts / Questions

Cauchy and Normal in several Dimensions ρ

ρ	1	2	3	4
$E_I(\mathbf{C})$	0	$-\ln 2 \approx -0.69$	-1	$-\ln 2 - \frac{1}{2} \approx -1.19$
$\sigma_I^2(\mathbf{C})$	$\frac{\pi^2}{4}$ ≈ 2.47	$\frac{\pi^2}{6}$ ≈ 1.64	$\frac{\pi^2}{4} - 1$ ≈ 1.47	$\frac{\pi^2}{6} - \frac{1}{4}$ ≈ 1.39
$E_I(\mathbf{N})$	$\frac{\gamma + \ln 2}{2}$ ≈ 0.64	$\frac{\gamma - \ln 2}{2}$ ≈ -0.06	$\frac{\gamma + \ln 2 - 1}{2}$ ≈ -0.36	$\frac{\gamma - \ln 2 - 1}{2}$ ≈ -0.56
$\sigma_I^2(\mathbf{N})$	$\pi^2/8$ ≈ 1.23	$\pi^2/24$ ≈ 0.41	$\pi^2/8 - 1$ ≈ 0.23	$\pi^2/24 - 1/4$ ≈ 0.16

Logarithmic Moments of other Multidimensional Distributions (Exponential, Uniform, Pareto,...) ?

Interesting Facts / Questions

'Simultaneous' Moments

For instance,

$$\int_0^1 \frac{(-\ln x)^{s-1} x^n}{1-x} dx = \Gamma(s) \zeta(s, n+1)$$

$$\int_0^1 \int_0^1 \frac{(-\ln xy)^{s-2} (xy)^n}{1-xy} dx dy = \Gamma(s) (\zeta(s, n) - n^{-s})$$

$$\int_0^1 \frac{(-\ln x)^{s-1} x^n}{1+x^2} dx = \frac{\Gamma(s)}{4^s} \left(\zeta\left(s, \frac{n+1}{4}\right) - \zeta\left(s, \frac{n+3}{4}\right) \right)$$

where $\zeta(s, m) = \sum_{k=0}^{\infty} \frac{1}{(k+m)^s}$ is the Hurwitz zeta function.

Note the shift '+m' in the denominator!

Applications / Outlook: Statistics

Traditional statistical modelling:

Data = Structure + Error

where the error is normally distributed, since $E(N) = 0$

Extension:

Data = Structure · Error

Here, a Cauchy error is a natural choice, since $E_I(C) = 0$

General Models

should employ both kinds of error

Applications / Outlook: Signal \leftrightarrow Noise

Signal vs. Noise Paradigm

- Convergence (towards a point),
e.g. of opinions, sedimentation
- Divergence (towards infinity),
e.g. polarization and confusion, erosion
- Middle Ground: Fractals and Distributions

Stochastic Distributions are a nice **compromise**:

- there are many distributions, and they are closely related
- centripetal 'force' \approx centrifugal 'force' (center vs. periphery)
- precise analysis: positive vs. negative expected information. . .
- they model Mandelbrot's States of Randomness:
Mild - Slow - Wild (e.g. the Normal vs. Power Laws)

Thank you!

Uwe Saint-Mont (2026): Information is Key.

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